

STA9715 - Test 2 - Formula Sheet

Inequalities

- If $\text{support}(X)$ is non-negative, $\mathbb{P}(X > x) \leq \mathbb{E}[X]/x$ (Markov)
- For any $X \sim (\mu, \sigma^2)$, $\mathbb{P}(|X - \mu| \geq k\sigma) \leq 1/k^2$ or $\mathbb{P}(|X - \mu| \geq k) \leq \sigma^2/k^2$ (Chebyshev)

Vector Arithmetic and Linear Algebra

- Vector addition $\mathbf{x} + \mathbf{y} = (x_1 + y_1, x_2 + y_2, \dots, x_n + y_n)$ Vector-times-scalar $\alpha\mathbf{x} = (\alpha x_1, \alpha x_2, \dots, \alpha x_n)$. (Both elementwise)
- Two-vector (dot / inner) product yields scalar: $\langle \mathbf{x}, \mathbf{y} \rangle = \mathbf{x} \cdot \mathbf{y} = x_1 y_1 + x_2 y_2 + \dots + x_n y_n$
- Vector norm: $\|\mathbf{x}\| = \sqrt{\sum_i x_i^2}$ - generalizes length or absolute value. Angle between vectors: $\cos \angle(\mathbf{x}, \mathbf{y}) = \langle \mathbf{x}, \mathbf{y} \rangle / \|\mathbf{x}\| \|\mathbf{y}\|$
- Matrix-vector multiplication: yields a vector: $\mathbf{A}\mathbf{x}$ element i is dot product of row i of \mathbf{A} with \mathbf{x} .
- Matrix-matrix multiplication: yields a matrix: $\mathbf{A}\mathbf{B}$ element (i, j) is dot product of row i of \mathbf{A} with column j of \mathbf{B} .
- Quadratic form: $\langle \mathbf{x}, \mathbf{A}\mathbf{x} \rangle = \|\mathbf{x}\|_{\mathbf{A}}^2 = \sum_{(i,j)} A_{ij} x_i x_j$. \mathbf{A} is *positive-definite* if all quadratic forms are positive (for $\mathbf{x} \neq \mathbf{0}$)
- Identity matrix \mathbf{I} is ones on diagonal; zeros elsewhere. $\mathbf{I}\mathbf{x} = \mathbf{x}$ and $\mathbf{A}\mathbf{I} = \mathbf{I}\mathbf{A} = \mathbf{A}$ for all \mathbf{x}, \mathbf{A}

Random Vectors

- Expectation is coordinate-wise: $\mathbb{E}[\mathbf{X}] = (\mathbb{E}[X_1], \mathbb{E}[X_2], \dots, \mathbb{E}[X_n])$
- Linear transforms: $\mathbb{E}[\mathbf{a} + \alpha\mathbf{X} + \beta\mathbf{Y}] = \mathbf{a} + \alpha\mathbb{E}[\mathbf{X}] + \beta\mathbb{E}[\mathbf{Y}]$ and $\mathbb{E}[\langle \mathbf{a}, \mathbf{X} \rangle] = \langle \mathbf{a}, \mathbb{E}[\mathbf{X}] \rangle$ Does not assume independence
- PDFs work via *multiple* integrals: $\mathbb{P}(\mathbf{X} \in A) = \iiint_A f_{\mathbf{X}}(\mathbf{x}) d\mathbf{x}$. CDFs are difficult
- If joint PDF factorizes $f_{(X,Y)}(x,y) = f_X(x)f_Y(y)$ then $X \perp Y$ (independence)
- Marginal PDF: $f_X(x) = \int_{-\infty, \infty} f_{(X,Y)}(x,y) dy$
- Conditional PDF: $f_{X|Y=y}(x) = f_{(X,Y)}(x,y)/f_Y(y)$. General form: $f_{X|Y \in A}(x) = \int_A f_{(X,Y)}(x,y) dy / \mathbb{P}(Y \in A)$

Covariance

- Covariance of two scalars: $\mathbb{C}[X, Y] = \mathbb{E}[(X - \mathbb{E}[X])(Y - \mathbb{E}[Y])] = \mathbb{E}[XY] - \mathbb{E}[X]\mathbb{E}[Y]$ (can be positive or negative)
- Self-covariance is variance: $\mathbb{C}[X, X] = \mathbb{V}[X]$
- Linear transforms: $\mathbb{C}[aX + b, cY + d] = ac\mathbb{C}[X, Y]$. For random vector \mathbf{X} and fixed matrix \mathbf{A} : $\mathbb{V}[\boldsymbol{\mu} + \mathbf{A}\mathbf{X}] = \mathbf{A}\mathbb{V}[\mathbf{X}]\mathbf{A}^T$.
- Correlation: $\rho_{X,Y} = \mathbb{C}[X, Y] / \sqrt{\mathbb{V}[X]\mathbb{V}[Y]}$
- Variance of a random vector is a (co)variance matrix: $\mathbb{V}[\mathbf{X}]_{ij} = \mathbb{C}[X_i, X_j]$
- Covariance quadratic forms give variance of linear combinations: $\mathbb{V}[\langle \mathbf{a}, \mathbf{X} \rangle] = \langle \mathbf{a}, \mathbb{V}[\mathbf{X}]\mathbf{a} \rangle = \sum_{ij} a_i a_j \mathbb{C}[X_i, X_j] \geq 0$
- Independence implies uncorrelated, but not the other way: $X \perp Y \implies \mathbb{C}[X, Y] = 0 \iff \mathbb{E}[XY] = \mathbb{E}[X]\mathbb{E}[Y]$

Normal Distribution

- Standard normal distribution. $Z \sim \mathcal{N}(0, 1)$. Mean Zero + Variance 1
- Standard normal PDF - $\phi(z) = \frac{1}{\sqrt{2\pi}} e^{-z^2/2}$. Standard normal CDF $\Phi(z) = \int_{-\infty}^z \phi(x) dx$ - no closed form.
- General normal distribution $X \sim \mathcal{N}(\mu, \sigma^2)$ - generated by scale+shift of standard normal $X \stackrel{d}{=} \mu + \sigma Z$.
- Normal PDF via standardization (z -score): $f_X(x) = \phi\left(\frac{x-\mu}{\sigma}\right) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-(x-\mu)^2/2\sigma^2}$. CDF: $\Phi\left(\frac{x-\mu}{\sigma}\right)$.
- Multivariate normal parameterized by mean vector and (co)variance matrix: $\mathbf{X} \sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$
- Standard multi-normal: $\mathbf{Z} \sim \mathcal{N}_n(\mathbf{0}_n, \mathbf{I}_n)$. PDF $f_{\mathbf{Z}}(\mathbf{z}) = (2\pi)^{-n/2} e^{-\|\mathbf{z}\|^2/2}$.
- General multi-normal $\mathbf{X} \stackrel{d}{=} \boldsymbol{\mu} + \boldsymbol{\Sigma}^{1/2}\mathbf{Z}$ where $\boldsymbol{\Sigma}^{1/2}$ is a matrix square root (Cholesky or symmetric).
- Bivariate normal PDF $f_{(X,Y)}(x,y) = \frac{1}{2\pi\sigma_X\sigma_Y\sqrt{1-\rho^2}} \exp\left(-\frac{1}{2[1-\rho^2]}\left[\left(\frac{x-\mu_X}{\sigma_X}\right)^2 - 2\rho\left(\frac{x-\mu_X}{\sigma_X}\right)\left(\frac{y-\mu_Y}{\sigma_Y}\right) + \left(\frac{y-\mu_Y}{\sigma_Y}\right)^2\right]\right)$
- Multivariate normal: any linear combination (weighted sum) of X_i is normal.
- If $\mathbb{C}[X_i, X_j] = 0$, then $X_i \perp X_j$ (for multi-normal, uncorrelated implies independent)
- If \mathbf{Z} is a standard normal n -vector, $\|\mathbf{Z}\|^2 = \sum_{i=1}^n Z_i^2$ has a χ^2 distribution with n degrees of freedom